

An analytical model for evaluating outage and handover probability of cellular wireless networks

Laurent Decreusefond, Philippe Martins, Thanh-Tung Vu

► **To cite this version:**

Laurent Decreusefond, Philippe Martins, Thanh-Tung Vu. An analytical model for evaluating outage and handover probability of cellular wireless networks. WPMC, Sep 2012, Taiwan. pp.1–5, 2012. <hal-00514212>

HAL Id: hal-00514212

<https://hal-imt.archives-ouvertes.fr/hal-00514212>

Submitted on 1 Sep 2010

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

An analytical model for evaluating outage and handover probability of cellular wireless networks

L. Decreusefond, P. Martins, T. T. Vu

Institut Telecom
Telecom Paristech
CNRS LTCI
Paris, France

Abstract—We consider stochastic cellular networks where base stations locations form a homogenous Poisson point process and each mobile is attached to the base station that provides the best mean signal power. The mobile is in outage if the SINR falls below some threshold. The handover decision has to be made if the mobile is in outage for some time slots. The outage probability and the handover probability is evaluated in taking into account the effect of path loss, shadowing, Rayleigh fast fading, frequency factor reuse and conventional beamforming. The main assumption is that the Rayleigh fast fading changes each time slot while other network components remain static during the period of study.

I. INTRODUCTION

In a wireless network, nodes can be modeled by fixed or stochastic pattern of points on the plane. Fixed points model can be finite or infinite and usually regular or lattice. This approach fails to capture the irregularity and randomness of a real network. For example, to model a wireless cellular network, the hexagonal cellular network is the model of this type most used. In reality, the base station (BS) nodes are usually fixed, it is not true that they are spatically periodical. Recently, stochastic model of nodes are more preferred. Node patterns can be represented by a stochastic process on the plane such as Poisson point process. It is worth to note that the stochastic models, although are more complicated at the first sight, usually lead to elegant and easy calculated formulas. In fact, all information obtained when studying both types of model are useful for the design or dimensioning processes of networks. In this paper we choose the stochastic approach and investigate a cellular network with homogeneous Poisson point process of BS.

However most of works relies on the assumption that a mobile once in the network is served by the nearest BS. This is due to considering path loss exponent model of radio propagation and remove the effect of fading. This assumption results to a so called Poisson-Voronoi cells model (for example, [2], page 63). Most works consider only the effect of fading but only the slow fading such as log-normal shadowing or fast fading such as Rayleigh fading. Besides, most of works consider the well known exponent propagation model. In this paper, the proposed model is slightly more general. Firstly, we consider a general model of path loss. Secondly, we are interested in a system spatically static but some temporal evolutionary elements. More precisely, we include both random general

slow fading and Rayleigh fading but that the slow fading being static in time, and the Rayleigh fading changes each time slot. Thirdly, once considering this, we make a very natural assumption that the mobile is served by the BS that provides the most strong mean signal power in time (best server). The mean signal power depends on path loss and slow fading. This choice of serving BS can be made either by the mobile or the operator. Thus, it can be though that our model is a generalization of Poisson-Voronoi cell model. If we assume that each BS generates an independent copy of a continuous shadowing random such as the one in [12], one can interpret continuous cell form. However if each slow fading field generated by a BS is independent fading fields then the cells can not be analytically identified, in particular they can not be measurable. We do not address this issue in this paper. Nevertheless we make assumption that the slow fading value of BSs to a mobile is independent, so including all above cases.

Once the mobile is served by one BS, the signal received by this BS will be the useful signal, and we assume that the considered system other signal received by other BS using the same frequency is interference. It is not true if we consider for example an advanced system in which the base stations are cooperative. However our model covers almost all existing cellular networks. To model the frequency reuse, we add a independent mark on our Poisson point process of BSs. A BS interferes other BSs that have the same mark. In addition to the interferences, the local noise can intervene. In order to make communication with the BS, the signal-to-noise-plus-interferences ratio (SINR) at this mobile location must be excess some threshold, in this case the mobile is covered, in contrary it is in outage. If the mobile is in outage during a period of time, i.e for some consecutive time slots, a handover decision has to be made. It can be made by the mobile, the served BS, the network system or even by a neighbor BS. In this paper, we are interested in the calculation of the outage probability and the handover probability in explicit forms. Since we assumes a homogenous Poisson point process of BSs, but not fixed patterns, these results does not depend on the position of the mobile and can be considered as global, meaning on all MS on the network.

This present paper benefits from results in the literature. In [9], Haenggi shows that the path loss fading process is Poisson

point process in real line in the case of path loss exponent model. In [1], [2] and [3], Baccelli and al. find analytical expressions for outage probability of networks where each node tries to connect with a destination of fixe distance or the nearest node in case of Rayleigh fading. In [6], Kelif and al. find a outage probability expression for cellular network by mean of the so-called fluid model. In [5], Ganti and al. find the interesting results about temporal and spatial correlation of wireless networks. In [10] and [11], outage probability of regular hexagonal cellular networks with reuse factor and adaptive beamforming is studied by simulation.

This paper is organized as following. In the section II we describe our model. In this section III we show that the path loss shadowing is a Poisson point process in real line. In the section IV we calculate outage probability. In the section V we calculate the handover probability. Section VI shows the numerical results and the difference between our model and the traditional hexagonal model.

II. SYSTEM MODEL AND SCENARIO

A. Propagation model

The signal radio propagation modelling is complicated, usually divided by a deterministic large scale path loss and the random fading components. The large scale path loss describes the channel at a microscopic level. If there are a BS (base station) located at y and an mobile located at x and the transmission power P , the mobile's received signal has the average power $L(y-x)P$ where L is the path loss function. We assume that L is measurable function on R^2 .

The most used path loss function is the path loss exponent law $L(z) = K|z|^{-\gamma}$ where $|z|$ refers to the Euclid norm of z . The parameter K depends on the frequency, the antenna height,... while the path loss exponent γ characterizes the environment under study. γ is typically in the range of (2,4), it may be greater if the environment is very dense urban. In fact, this path loss model is not correct for small distances and has infinite mean of interference for Poisson patterns of BSs [1]. To avoid theses problems, one can use the modified path loss exponent model $L(z) = K(\max\{R_0, |z|\})^{-\gamma}$ where R_0 is a reference distance.

In addition to the deterministic large scale effect, there are two random factors can be considered. The first, called *shadowing* or *shadow fading*, represents the signal attenuation caused by a large obstacle such as building. The second, called *fast fading*, represents the impact of multipath phenomena, or in other word many objects scatter the signal. The shadowing can be considered as constant during a period of communication of the mobile while the fast fading changes each time slot. If there is no beamforming technique is used, the received signal power from BS y to MS x at the time slot l will be

$$P_{yx}[l] = r_{y,x}[l]h_{yx}L(y-x)P, \quad (1)$$

where $\{h_{yx}\}_{x,y \in R^2}$ are copies of a random variable H while $\{r_{yx}[l]\}$ are independent copies of R which is an exponential random variable of mean $1/\mu$. We suppose that for each x , the random variables h_{yx} for all $y \in R^2$ are independent. We

define p_H the probability density function of H and $F_H(\beta) = P(H \geq \beta) = \int_{\beta}^{\infty} p_H(t)dt$. The most used shadowing random model is log-normal shadowing, for which H is a log-normal random variable. In this case, we can write $H \sim 10^{G/10}$ where $G \sim \mathcal{N}(0, \sigma^2)$.

In short word, we suppose that during the period of study, the shadowing remains constant while the Rayleigh fast fading changes at each time slot.

Consider the path loss shadowing process $\Xi = \{\xi_i = (h_{yx}L(y-x)P)^{-1}\}$. We make the following assumptions:

Assumption 1: For each $x \in R^2$, $\{h_{yx}\}_{y \in R^2}$ are independent.

Assumption 2: H admits a continuous probability density function on $(0, \infty)$.

Assumption 3: Define $B(\beta) = \int_{R^2} F_H((L(z)P\beta)^{-1})dz$, then $0 < B(\beta) < \infty$ for all $\beta > 0$.

By the displacement theorem we will show that Ξ is a simple Poisson point process in the real line $(0, +\infty)$ (proposition 1) of intensity $\Lambda(dt) = \lambda_B B'(\beta) > 0$. Hence we have the right to reorder Ξ such that $\xi_0 < \xi_1 < \dots$ and for simplicity we do it.

B. Poisson point process of BSs

We assume the homogenous Poisson point process of BSs $\Pi_B = \{y_0, y_1, \dots\}$ of intensity λ_B on R^2 , each BS transmit has constant transmitted power P . For any details on Poisson point process we refer to [2].

Assumption 4: Once being in the network the mobile x is firstly attached to (or served by) the BSs that provide the best *average* signal strength in *time*. In other word, it is attached by y_0 (after reordering and renumbering Ξ).

C. Beamforming model

We consider the conventional beamformer technique with n_t antennas. The power radiation pattern for a conventional beamformer is a product of array factor and radiation pattern of a single antenna. If ϕ is the look direction (toward which the beam is steered), the array gain in the direction θ is given by ([11], [10]):

$$\frac{\sin^2(n_t \frac{\pi}{2} (\sin(\theta) - \sin(\phi)))}{n_t \sin^2(\frac{\pi}{2} (\sin(\theta) - \sin(\phi)))} g(\theta),$$

where $g(\theta)$ is the gain in the direction θ with one antenna. For simplicity we assume that the BS always steers to the direction of the serving MS and the gain $g(\theta)$ is positive constant on $(-\pi/2, \pi/2)$ and 0 otherwise (zero front-to-back power ratio). Hence, the interference signal power from a BS to a MS attached by an other BS using the same frequency in the direction θ will be reduced by a factor of:

$$a(\theta) = 1_{\{\theta \in (-\pi/2, \pi/2)\}} \frac{\sin^2(n_t \frac{\pi}{2} (\sin(\theta)))}{n_t^2 \sin^2(\frac{\pi}{2} (\sin(\theta)))}$$

If the beamforming technique is not used, we will simply use $a(\theta) = 1$.

D. Frequency reuse

We add a mark to each BSs e_i . The marks e_i are independent copies of the random variable E who is uniformly distributed on $\{1, 2, \dots, k\}$ where k will be called the frequency reuse factor. The BSs that have the same mark interfere between themselves. Our reuse model can be considered as the worst case where the bandwidth is divided into k subband and each BS is randomly attributed a sub band. It is contrast to the hexagonal network pattern where the interfering BS must be placed far from a reference BS.

E. SINR

Assume that each other BS using the same frequency is always serving a MS, and the MS x is in the direction θ_i which is i.i.d chosen on $(-\pi, \pi)$ of the BS i . The SINR at the time slot l is defined as:

$$s_x[l] = \frac{r_{y_0x}[l]\xi_0^{-1}}{N + \sum_{i \neq 0} 1_{\{e_i=e_0\}} a(\theta_i) r_{y_ix}[l]\xi_i^{-1}}, \quad (2)$$

where N is a constant noise power. The term $I_x = \sum_{i \neq b(x)} 1_{\{e_i=e_0\}} a(\theta_i) r_{y_ix}[l]\xi_i^{-1}$ is the sum of all interferences. In order to make communication with the attached BS, the SINR must not fall below some threshold T .

F. Handover decision

We consider a simple SINR based decision. The handover should be made if the MS is in outage for n consecutive time slots.

G. Scenario

The scenario is as following:

- Realization of a snapshot of BSs y_i , slow fading h_{y_ix} and the frequency e_i .
- Attachment of mobile x to best BS, (y_0 after reordering Ξ).
- Realization of the directions θ_i for interfering BS y_i .
- At time slot l , realization of Rayleigh fast fading $r_{y_ix}[l]$ and calculate the SINR $s_x[l]$. If $s_x[l] < T$ then the mobile is in outage otherwise it is covered.
- If the mobile is in outage for n consecutive time slots then the handover should be made.

The outage probability is then $p_o(T) = P(s_x[l] < T)$ and the handover decision probability is $p_{ho}(T) = P(s_x[l] < T, \dots, s_x[l+n-1] < T)$. We also define the coverage probability $p_C(T) = P(s_x[l] \geq T)$.

H. Interference limited case

We are particularly interested in the interference limited regime when the noise power N is negligible or nearly equal to zero as it happens usually in a real network. We can set $N = 0$. The outage probability calculated in the interference regime can be considered as an upper bound for the outage probability in the general case.

III. POISSON POINT PROCESS OF PATH LOSS SHADOWING

A. General case

Proposition 1: Ξ is a Poisson point process on $R^+ = (0, \infty)$ with intensity density $\Lambda(dt) = \lambda_B B'(t) dt$.

Proof: Define the marked point process $\Pi^x = \{y_i, h_{y_ix}\}_{i=0}^\infty$. It is a Poisson point process of intensity $\lambda_B dy \otimes f_H(t) dt$ because the marks are i.i.d. We consider the probability kernel $p((z, t), A) = 1_{\{(L(z)Pt)^{-1} \in A\}}$ for all Borel $A \in R^+$ and apply the displacement theorem ([2], theorem 1.3.9) to obtain that the point process Ξ is Poisson point process of intensity

$$\Lambda(A) = \lambda_B \int_{R^2 \otimes R} 1_{\{(L(z)tP)^{-1} \in A\}} p_H(t) dz dt.$$

We now show that $\Lambda([0, \beta]) = \lambda_B B(\beta)$. Indeed,

$$\begin{aligned} \Lambda([0, \beta]) &= \lambda_B \int_{R^2 \otimes R} 1_{\{t \geq (\beta P L(z) P)^{-1}\}} p_H(t) dz dt \\ &= \lambda_B \int_{R^2} F_H((\beta L(z) P)^{-1}) dz \\ &= \lambda_B B(\beta). \end{aligned}$$

Finally $B(\beta)$ admits a derivative:

$$B'(\beta) = \beta^{-2} \int_{R^2} \frac{1}{L(z)P} p_H((\beta L(z)P)^{-1}) dz. \quad (3)$$

This concludes the proof. \blacksquare

The CDF and PDF of ξ_m are easily derived according to the property of Poisson point process:

Lemma 1: The complementary cumulative distribution function of ξ_m is given by:

$$P(\xi_m > t) = e^{-\lambda_B B(t)} \sum_{i=0}^m \frac{(\lambda_B B(t))^i}{i!}, \quad (4)$$

and its probability density function is given by

$$p_{\xi_m}(t) = \lambda_B^{m+1} B'(t) B(t)^m e^{-\lambda_B B(t)}. \quad (5)$$

Proof: The event " $\xi_m > t$ " is equivalent to the event "in the interval $[0, t]$ there is at most m points" and the number of points in this interval follows a Poisson random variable of mean $\lambda_B B(t)$, so:

$$P(\xi_m < t) = e^{-\lambda_B B(t)} \sum_{i=0}^m (\lambda_B B(t))^i$$

The PDF is thus given by $p_{\xi_m}(t) = -\frac{d}{dt} P(\xi_m < t)$, and after some simple manipulations we obtain the equation (5) \blacksquare

B. Special cases

In this section we derive closed forms for $B(\beta)$ in some special cases.

a) Path loss exponent model:

Lemma 2: If $L(z) = K|z|^{-\gamma}$ then:

$$B(\beta) = C_1 \beta^{\frac{2}{\gamma}}, \quad (6)$$

where $C = \pi(PK)^{\frac{2}{\gamma}} E(H^{\frac{2}{\gamma}})$.

Proof: We have:

$$\begin{aligned} B(\beta) &= 2\pi \int_0^\infty r \mathbf{1}_{\{tPK\beta \geq r^\gamma\}} p_H(t) dr dt \\ &= 2\pi \int_0^\infty p_H(t) dt \int_0^{(tKP\beta)^{1/\gamma}} r dr \\ &= \pi(PK)^{\frac{2}{\gamma}} \beta^{\frac{2}{\gamma}} \int_0^\infty p_H(t) t^{\frac{2}{\gamma}} dt \\ &= \pi(PK)^{\frac{2}{\gamma}} E(H^{\frac{2}{\gamma}}) \beta^{\frac{2}{\gamma}}. \end{aligned}$$

Hence the result. \blacksquare

Remark that this result can be derived from [9]. We observe that the distribution of the point process Ξ does depend only on $E(H^{\frac{2}{\gamma}})$ but not on the distribution of shadowing H itself. This phenomenon can be explained as in [4](page 159).

b) Modified path loss exponent model:

Lemma 3: If $L(z) = K(\max\{R_0, |z|\})^{-\gamma}$ then:

$$B(\beta) = C_1 \beta^{\frac{2}{\gamma}} \int_{\frac{R_0^\gamma}{\beta PK}}^\infty t^{\frac{2}{\gamma}} p_H(t) dt, \quad (7)$$

where $C_1 = \pi(PK)^{\frac{2}{\gamma}}$. In addition, we have:

$$B'(\beta) = \frac{2}{\gamma} \beta^{-1} B(\beta) + \pi R_0^2 p_H\left(\frac{R_0^\gamma}{PK\beta}\right). \quad (8)$$

If the slow fading is lognormal shadowing $H \sim 10^{G/10}$ where $G \sim \mathcal{N}(0, \sigma^2)$ we have:

$$B(\beta) = C_1 \beta^{\frac{2}{\gamma}} e^{\frac{2\sigma_1^2}{\gamma}} Q\left(\frac{-\ln \beta - \ln(PK R_0^{-\gamma})}{\sigma_1} - \frac{2\sigma_1}{\gamma}\right) \quad (9)$$

where $Q(u) = \frac{1}{\sqrt{2\pi}} \int_u^\infty e^{-u^2/2} du$ is the Q -function and $\sigma_1 = \frac{\sigma \ln 10}{10}$.

Proof: Similarly to the pathloss exponent model case, we have:

$$\begin{aligned} B(\beta) &= 2\pi \int_{R^2} r F_H((\max\{R_0, r\})^{-\gamma} (PK\beta)^{-1}) dr \\ &= \pi R_0^2 F_H(R_0^\gamma (PK\beta)^{-1}) + \\ &\quad + 2\pi \int_{\frac{R_0^\gamma}{\beta PK}}^\infty p_H(t) dt \int_{R_0}^{(tKP\beta)^{1/\gamma}} r dr \\ &= C_1 \beta^{\frac{2}{\gamma}} \int_{\frac{R_0^\gamma}{PK\beta}}^\infty t^{\frac{2}{\gamma}} p_H(t) dt. \end{aligned}$$

We obtain the equation (7). Derivative two sides of that equation and do some simple manipulations we obtain the

equation (8). In the case of lognormal shadowing we have:

$$\begin{aligned} B(\beta) &= C_1 \beta^{\frac{2}{\gamma}} \int_{\frac{R_0^\gamma}{PK\beta}}^\infty \frac{1}{\sqrt{2\pi\sigma_1^2} t} t^{\frac{2}{\gamma}} e^{-\frac{(\ln t)^2}{2\sigma_1^2}} dt \\ &= C_1 \beta^{\frac{2}{\gamma}} \int_{\ln \frac{R_0^\gamma}{PK\beta}}^\infty \frac{1}{\sqrt{2\pi\sigma_1^2}} e^{\frac{2u}{\gamma}} e^{-\frac{u^2}{2\sigma_1^2}} du \\ &= C_1 \beta^{\frac{2}{\gamma}} e^{\frac{2\sigma_1}{\gamma}} \int_{\ln \frac{R_0^\gamma}{PK\beta}}^\infty \frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{(u - \frac{2\sigma_1}{\gamma})^2}{2\sigma_1^2}} du. \end{aligned}$$

Here the results. \blacksquare

As a consequence, both the exponent path loss model and its modified model satisfy the assumption 3.

IV. OUTAGE ANALYSIS

A. General case

Here, we remark that the outage probability and the coverage probability do not depend on the time index. So we can drop the time slot parameter in this section. The expression of the SINR can be rewritten as:

$$s_x = \frac{r_{y_0 x} \xi_0^{-1}}{N + \sum_{i \neq 0} \mathbf{1}_{\{e_i = e_0\}} a(\theta_i) r_{y_i x} \xi_i^{-1}}. \quad (10)$$

The outage probability is calculated as below:

Theorem 1: The outage probability is given by

$$p_o = 1 - \lambda_B \int_0^\infty B'(\beta) e^{-\lambda_B B(\beta) - NT\mu\beta - \frac{\lambda_B}{2\pi k} D(\beta)} d\beta$$

where $D(\beta) = \int_{-\pi}^\pi d\theta \int_\beta^\infty B'(\xi) \frac{d\xi}{1 + \xi(T\beta a(\theta))^{-1}}$.

Proof: To calculate the outage probability $P(s_x < T)$, we will calculate the coverage probability $P(s_x \geq T)$.

We first consider the conditional probability $P(s_x \geq T | \xi_0 = \beta)$. Because $r_{y_0 x}$ is an exponential random variable of mean $1/\mu$ we have:

$$\begin{aligned} P(s_x \geq T | \xi_0 = \beta) &= P(r_{y_0 x} \geq T\beta(N + I_x(\beta)) | \xi_0 = \beta) \\ &= E(e^{-\mu T\beta(N + I_x(\beta))} | \xi_0 = \beta) \\ &= e^{-NT\mu\beta} \mathcal{L}_{I_x(\beta)}(T\mu\beta) \end{aligned}$$

where $I_x(\beta)$ is the distribution of the random variable I_x given on the event $(\xi_0 = \beta)$ and $\mathcal{L}_{I_x(\beta)}$ is its Laplace transform. Conditioning on the event $(\xi_0 = \beta)$ the point process $\{\xi_i\}_{i>0}$ is a Poisson point process on (β, ∞) with intensity $\lambda_B B'(\xi) d\xi$ according to the strong Markov property. By thinning theorem, the point process $\{\xi_i\}_{i>0, e_i=e_0}$ is a Poisson point process on $(0, \beta)$ with intensity $\frac{\lambda_B}{k} B'(\xi) d\xi$. Hence, $\mathcal{L}_{I_x(\beta)}$ can be calculated as follows ([2], shot noise theory):

$$\begin{aligned} \mathcal{L}_{I_x(\beta)}(u) &= e^{-\int_\beta^\infty \frac{\lambda_B}{2\pi k} B'(\xi) (1 - E(e^{-a(\theta)u\xi^{-1}R}) d\xi} \\ &= e^{-\frac{\lambda_B}{2\pi k} \int_\beta^\infty B'(\xi) d\xi \int_0^\infty dr \int_{-\pi}^\pi \mu e^{-\mu r} (1 - e^{-a(\theta)ur\xi^{-1}}) d\theta} \\ &= e^{-\frac{\lambda_B}{2\pi k} \int_{-\pi}^\pi d\theta \int_\beta^\infty B'(\xi) \frac{d\xi}{1 + \xi\mu(a(\theta))^{-1}}}. \end{aligned}$$

We get that:

$$\begin{aligned} P(s_x \geq T | \xi_0 = \beta) &= \\ &= e^{-NT\mu\beta - \frac{\lambda_B}{2\pi k} \int_{-\pi}^\pi d\theta \int_\beta^\infty B'(\xi) \frac{d\xi}{1 + \xi(T\beta a(\theta))^{-1}}}, \end{aligned}$$

thus

$$= e^{-NT\mu\beta - \frac{\lambda_B}{2\pi k} D(\beta)}. \quad (11)$$

Since the distribution density of ξ_0 is $\lambda_B B'(\beta) e^{-\lambda_B B(\beta)}$ (proposition 1), by averaging over all ξ_0 we obtain the equation (11). ■

B. Special cases

c) Interference limited:

Collary 1: In the interference-limited regime, we have

$$p_o(T) = 1 - \lambda_B \int_0^\infty B'(\beta) e^{-\lambda_B B(\beta) - \frac{\lambda_B}{2\pi k} D(\beta)} d\beta. \quad (12)$$

d) Path loss exponent model:

Collary 2: If $L(z) = K|z|^{-\gamma}$ we have:

$$p_o(T) = 1 - \int_0^\infty e^{-M\alpha - G\alpha^{\frac{\gamma}{2}}} d\alpha \quad (13)$$

where $M := M(k, T, \gamma) = 1 + \frac{1}{2\pi k} \int_{-\pi}^\pi d\theta \int_1^\infty \frac{du}{1 + (Ta(\theta))^{-1} u^{\frac{\gamma}{2}}}$ and $G = NT\mu(\lambda_B C)^{-\frac{\gamma}{2}}$.

Proof: Since $B(\xi) = C.\xi^{\frac{2}{\gamma}}$ and $B'(\xi) = \frac{2C}{\gamma}\xi^{\frac{2}{\gamma}-1}$ we have:

$$\begin{aligned} D(\beta) &= \int_{-\pi}^\pi d\xi \int_\beta^\infty \frac{2C}{\gamma} \xi^{\frac{2}{\gamma}-1} \frac{d\theta}{1 + \xi(T\beta a(\theta))^{-1}} \\ &= C.\beta^{\frac{2}{\gamma}} \int_{-\pi}^\pi d\theta \int_\beta^\infty \frac{d(\frac{\xi}{\beta})^{\frac{2}{\gamma}}}{1 + \frac{\xi}{\beta}(Ta(\theta))^{-1}} \\ &= C.\beta^{\frac{2}{\gamma}} \int_{-\pi}^\pi d\theta \int_1^\infty \frac{du}{1 + (Ta(\theta))^{-1} u^{\frac{\gamma}{2}}}. \end{aligned}$$

Plug it into (11) we have:

$$\begin{aligned} p_c(T) &= \int_0^\infty \frac{2\lambda_B C}{\gamma} \beta^{\frac{2}{\gamma}-1} e^{-\lambda_B C M \beta^{\frac{2}{\gamma}} - NT\mu\beta} d\beta \\ &= \int_0^\infty e^{-M\alpha - G\alpha^{\frac{\gamma}{2}}} d\alpha. \end{aligned}$$

Remark that if $\gamma = 4$, we can find that:

$$\begin{aligned} M &= 1 + \frac{1}{2\pi k} \int_{-\pi}^\pi d\theta \int_1^\infty \frac{du}{1 + (Ta(\theta))^{-1} u^2} \\ &= 1 + \frac{1}{2\pi k} \int_{-\pi}^\pi \sqrt{Ta(\theta)} \left(\frac{\pi}{2} - \arctan \frac{1}{\sqrt{Ta(\theta)}} \right) d\theta \end{aligned}$$

and

$$\begin{aligned} p_o(T) &= 1 - e^{\frac{M^2}{4G}} \int_0^\infty e^{-(\sqrt{G}\alpha + \frac{M}{2\sqrt{G}})^2} d\alpha \\ &= 1 - \frac{\sqrt{2\pi}}{G} e^{\frac{M^2}{4G}} Q\left(\frac{M}{2\sqrt{G}}\right). \end{aligned}$$

e) Interference limited and path loss exponent model:

In this case, the outage probability is easily derived from (13) by setting $N = 0$.

Collary 3: If $L(z) = K|z|^{-\gamma}$ and $N = 0$ we have:

$$p_o(T) = 1 - \frac{1}{M}. \quad (14)$$

C. Observations and interpretations

Some interesting facts are observed from above results:

- Rewrite the expression of SINR as

$$s_x[l] = \frac{\bar{r}_{y_{0x}}[l] \xi_0^{-1}}{\mu N + \sum_{i \neq 0} 1_{\{e_i = e_0\}} a(\theta_i) \bar{r}_{y_{ix}}[l] \xi_i^{-1}}$$

where $\bar{r}_{y_{0x}}[l] = \mu r_{y_{0x}}[l]$. Since $r_{y_{0x}}[l]$ is an exponential random variable of mean $1/\mu$, $\bar{r}_{y_{0x}}[l]$ is an exponential random variable of mean 1. Hence by the above equation it is expected that the outage probability depends on the product μN but not directly on μ and N . It is increasing function of $N\mu$ which is confirmed by the equation (11). The fact that the outage probability is the increasing function of μ and N is quite natural, the increase of noise or the degrade of the channel fast fading always makes the system work worse.

- It is also expected that in the interference limited case ($N = 0$) the outage probability does not depend on μ . It is confirmed by the equation (12). Physically it means that in the absence of noise, the fast fading increases or degrades the channels to the MS of the serving BS and the interfering BS at the same level, thus the SINR will not change.
- In the interference limited and exponent path loss model case, the outage probability does not depend on μ , the BS density λ_B , or the distribution of shadowing H . It is due to the scaling property of the exponent path loss and the homogeneous Poisson point process. The outage probability is a decreasing function of the path loss exponent γ , reflecting the fact that bad propagation environment degrades the received SINR.
- In the presence of noise $N > 0$ and exponent path loss model case, the outage probability is a increasing function of λ_B . Hence, it can be thought that the more an operator installs BSs, the better the network is. In addition, if the density of BSs goes to infinite then outage will never occur. However it is not true. In fact, if the density of BSs is very high, the distance between a MS and its serving BS and some interfering BSs is relatively close. Here, the exponent path loss model is no longer valid since it is not accurate at small distance. If the modified exponent path loss is used which is more appreciate, the outage probability must converge to 0. The outage probability is also a increasing function of $E(H^{\frac{2}{\gamma}})$, and if the shadowing H follows lognormal distribution then the outage probability will be increasing function of σ . We recover an other well known fact: the increase of uncertainty of the radio channel degrades the performance of the network.

V. HANDOVER ANALYSIS

A. General case

If the MS is in outage in n consecutive time slots, a handover decision has to be made. Keep in mind that only the Rayleigh fast fading changes each time slot, and the other

network components do not change. Let A_l be the event that the mobile being in outage in the time slot l , and A_l^c its complement and observe that in fact $P(\cap_{i=1}^m A_{j_i}^c) = P(\cap_{i=1}^m A_i^c)$. By definition $p_{ho} := P(\cap_{i=1}^n A_{l+i-1}) = P(\cap_{i=1}^n A_i)$. We have

$$\begin{aligned} p_{ho} &= 1 + \sum_{m=1}^n (-1)^m \sum_{j_1 \neq \dots \neq j_m \in \{1, \dots, n\}} P(\cap_{i=1}^m A_{j_i}^c) \\ &= 1 + \sum_{m=1}^n (-1)^m \frac{n!}{m!(n-m)!} P(\cap_{i=1}^m A_i^c). \end{aligned}$$

Theorem 2: The handover probability is given by:

$$p_{ho} = 1 + \sum_{m=1}^n (-1)^m \frac{n!}{m!(n-m)!} q_m,$$

where $q_m = P(\cap_{i=1}^m A_i^c)$ is given by:

$$q_m = \int_0^\infty \lambda_B B'(\beta) e^{-\lambda_B B(\beta) - NT\mu\beta - \frac{\lambda_B}{2\pi k} D_m(\beta)} d\beta,$$

and $D_m(\beta) = \int_{-\pi}^\pi d\theta \int_\beta^\infty B'(\xi) (1 - (\frac{1}{1+T\beta a(\theta)\xi^{-1}})^m) d\xi$.

Proof:

We need to calculate the probability $P(\cap_{i=1}^m A_i^c)$ that is the probability that the mobile is covered in m different time slots. The calculation is similar the that in section IV. We begin with calculating the conditional probability $P(\cap_{i=1}^m A_i^c | \xi_0 = \beta)$:

$$\begin{aligned} P(\cap_{i=1}^m A_i^c | \xi_0 = \beta) &= P(s_x[1] \geq T, \dots, s_x[m] \geq T | \xi_0 = \beta) \\ &= P(r_{y_0x}[i] \geq \beta(TN + I_x[i]) | i = 1..m | \xi_0 = \beta) \\ &= E(e^{-\mu(mTN\beta + \sum_{i=1}^m I_x(\beta)[i])} | \xi_0 = \beta) \\ &= e^{-mNT\mu\beta} \mathcal{L}_{\sum_{i=1}^m I_x(\beta)[i]}(T\mu\beta) \end{aligned}$$

where $I_x(\beta)[i]$ is the distribution of the random variable $I_x[i]$ given $(\xi_0 = \beta)$. We have :

$$\sum_{i=1}^m I_x(\beta)[i] = \sum_{j=1}^\infty 1_{\{e_i = e_0\}} \xi_i^{-1} a(\theta_i) (\sum_{i=1}^m r_{y_{ix}}[i]).$$

As the random variables $r_{y_{ix}}[i]$ are independent copies of the exponential random variable R , the random variables $\sum_{i=1}^m r_{y_{ix}}[i]$ are also i.i.d and the common Laplace transform of the later are :

$$\begin{aligned} \mathcal{L}_{\sum_{i=1}^m r_{y_{ix}}[i]}(u) &= (\mathcal{L}_R(u))^m \\ &= \left(\frac{\mu}{\mu + u}\right)^m. \end{aligned}$$

The Laplace transform of $\sum_{i=1}^m I_x(\beta)[i]$ is now:

$$\mathcal{L}_{\sum_{i=1}^m I_x(\beta)[i]}(u) = e^{-\frac{\lambda_B}{2\pi k} \int_{-\pi}^\pi d\theta \int_\beta^\infty B'(\xi) (1 - (\frac{\mu}{\mu + a(\theta)\xi^{-1}u})^m) d\xi}.$$

The conditional probability is then given by:

$$P(\cap_{i=1}^m A_i^c | \xi_0 = \beta) = e^{-mNT\mu\beta - \frac{\lambda_B}{2\pi k} D_m(x)}.$$

By averaging with respect to ξ_0 , we have:

$$q_m = \int_0^\infty \lambda_B B'(\beta) e^{-\lambda_B B(\beta) - NT\mu\beta - \frac{\lambda_B}{2\pi k} D_m(\beta)} d\beta.$$

This concludes the proof.

B. Special cases

We can obtain more closed expression for q_m in some special cases.

f) *Interference limited:*

Collary 4: In the interference limited regime $N = 0$, we have:

$$q_m = \int_0^\infty \lambda_B B'(\beta) e^{-\lambda_B B(\beta) - \frac{\lambda_B}{2\pi k} D_m(\beta)} d\beta.$$

g) *Path loss exponent model:*

Collary 5: If $L(z) = K|z|^{-\gamma}$ then:

$$q_m = \int_0^\infty e^{-M_m \alpha - G \alpha^{\frac{\gamma}{2}}} d\alpha$$

where $M_m = 1 + \frac{1}{2\pi k} \int_{-\pi}^\pi d\theta \int_1^\infty (1 - (\frac{1}{1+T a(\theta) u^{-\frac{\gamma}{2}}})^m) du$.

Proof: The proof follows the same lines as the proof of 2 ■

Closed expression of q_m is obtained in the case $\gamma = 4$:

$$q_m = \frac{\sqrt{2\pi}}{G} e^{\frac{M_m^2}{4G}} Q\left(\frac{M_m}{2\sqrt{G}}\right).$$

h) *Interference limited and path loss exponent model:*

Collary 6: If $N = 0$ and $L(z) = K|z|^{-\gamma}$ we have:

$$q_m = \frac{1}{M_m}.$$

C. Observations and interpretations

Some interesting facts are observed from above results and they are similar to the properties of outage probability:

- The handover probability is increasing function of $N\mu$.
- In the interference limited and exponent path loss model case, the handover probability does not depend on μ , the BSs density λ_B , nor the distribution of shadowing H . The handover probability is a decreasing function of the path loss exponent γ .
- In the presence of noise $N > 0$ and exponent path loss model case, the handover probability is a increasing function of λ_B . Thus, in this case the more an operator installs BSs, the less a MS has to do handover. But it is not true in a real system. As previously explained, in the case of very dense BSs, the pathloss exponent model is no longer accurate for small distance. The handover probability is also a increasing function of $E(H^{\frac{2}{\gamma}})$, therefore if the shadowing H is lognormal shadowing the handover probability will be increasing function of σ .

VI. NUMERICAL RESULTS AND COMPARISON TO THE HEXAGONAL MODEL

We place a MS at the origin o and consider a region $B(o, R_g)$ where $R_g = 10.000(m)$. The BSs are distributed as a Poisson point process in this region. The path loss exponent model is considered. The default values of model are placed on the table I. They are not changed throughout the simulation. ■

TABLE I
MODEL PARAMETERS' DEFAULT VALUES

K	P	n_t	μ
-20 (dB)	0(dBm)	8	1

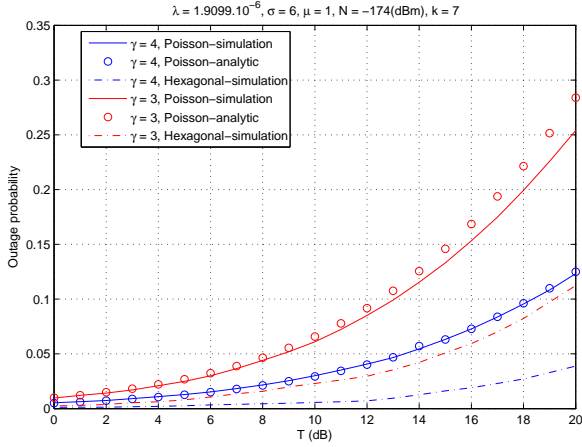


Fig. 1. Outage probability vs SINR threshold

In literature, the hexagonal model is widely used and studied so we would like to compare two models. For a fair comparison, the density of BSs must be chosen to be the same, i.e the area of a hexagonal cell will be $1/\lambda_B$. Unlike the Poisson model where each BS is randomly assigned a frequency, in the hexagonal model, the frequencies are well assigned so that an interfering BS is far from the transmitting BS and BSs of different frequency are grouped in reuse patterns. The reuse factor k in the hexagonal model is determined by $k = i^2 + j^2 + ij$ where integers i, j are the relative location of co-channel cell. The MS is uniformly chosen on the surface of the center cell. The same signal propagation model and the scenario described in the section II-G are applied in the hexagonal model.

The figure 1 shows the outage probability versus the SINR threshold of the Poisson model and the hexagonal model in the case $k = 7$. As we can see, the outage probability in the case of Poisson model is always greater than that of hexagonal model which is intuitive. The different is about 8 (dB) in the case $\gamma = 4$ and 6(dB) in the case $\gamma = 3$.

In the figure 3 plotted the outage probability as a function of γ . We can see that the outage probability is a decreasing function of γ as theoretically observed. In the figure 4 plotted the handover probability versus the SINR threshold of Poisson model and hexagonal model. If the reuse factor k increase, the MS has to do less handover. Thus, increase the reuse factor has a positive effect on the system performance not only in term of outage but also in term of handover.

VII. CONCLUSION

In this paper we have investigated the outage and handover probabilities of wireless cellular networks taking into account the reuse factor, the beamforming, the path loss, the slow

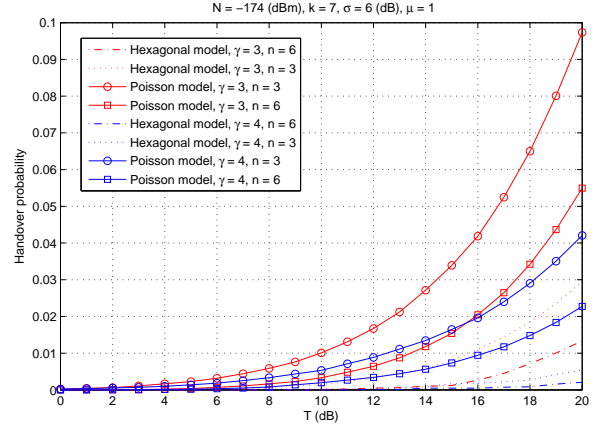


Fig. 2. Handover probability vs SINR threshold

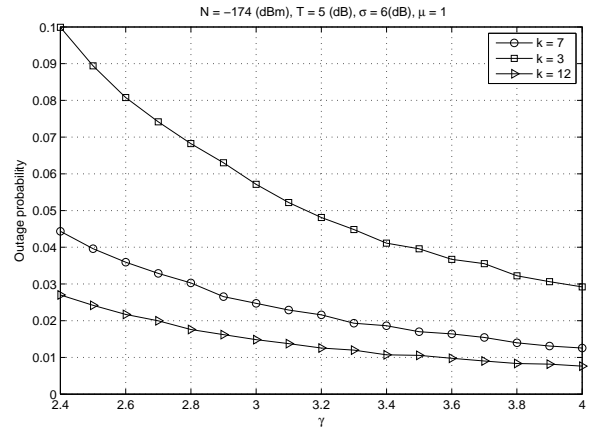


Fig. 3. Outage probability vs path loss exponent γ , Poisson model

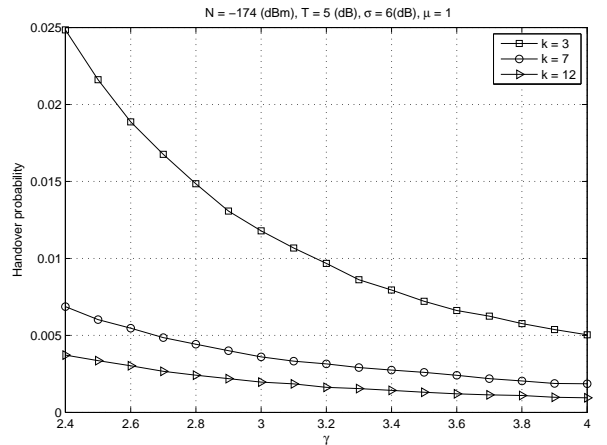


Fig. 4. Handover probability vs path loss exponent γ , Poisson model, $n = 3$

fading and the fast fading. We valid our model by simulation and compare numerical results to that of hexagonal model. The analytical expressions derived in the this paper can be considered as an upper bound for a real system.

REFERENCES

- [1] F. Baccelli, B. Blaszczyzyn and P. Muhlethaler , An Aloha protocol for multihop mobile wireless networks. *IEEE Trans. Inf. Theory* 52, 421-436, 2006
- [2] F. Baccelli and B. Blaszczyzyn , *Stochastic Geometry and Wireless Networks, Volume I - Theory* vol. 3, No 3-4 of *Foundations and Trends in Networking*, NoW Publishers, 2009
- [3] F. Baccelli and B. Blaszczyzyn , *Stochastic Geometry and Wireless Networks, Volume II - Applications* vol. 4, No 1-2 of *Foundations and Trends in Networking*, NoW Publishers, 2009
- [4] M. Haenggi and R. Krishna Ganti, *Interference in Large Wireless Networks*, Vol. 3, No. 2 of *Foundations and Trends in Networking*, NoW Publishers, 2009
- [5] R. K. Ganti and M. Haenggi, Spatial and Temporal Correlation of the Interference in ALOHA Ad Hoc Networks, *IEEE Communications Letters*, vol. 13, pp. 631-633, Sept. 2009
- [6] J. M. Kelif, M. Coupechoux and Ph. Godlewski, Spatial Outage Probability for Cellular Networks, *IEEE Globecom*, 2007
- [7] V.I. Mordachev, S. Loyka, On Node Density Outage Probability Tradeoff in Wireless Networks, *IEEE Journal on Selected Areas in Communications* (Special Issue on Stochastic Geometry and Random Graphs for the Analysis and Design of Wireless Networks), v. 27, N. 7, pp. 1120-1131, Sep. 2009
- [8] M. Haenggi, J. G. Andrews, F. Baccelli, O. Dousse, and M. Franceschetti, *Stochastic Geometry and Random Graphs for the Analysis and Design of Wireless Networks*, *IEEE Journal on Selected Areas in Communications*, vol. 27, pp. 1029-1046, Sept. 2009. Invited Paper
- [9] M. Haenggi, A Geometric Interpretation of Fading in Wireless Networks: Theory and Applications, *IEEE Trans. on Information Theory*, vol. 54, pp. 5500-5510, Dec. 2008
- [10] X. Lagrange, CIR cumulative distribution in a regular network, internal research report, ENST Paris, 2000
- [11] M. Maqbool, M. Coupechoux and Ph. Godlewski, Comparison of Various Frequency Reuse Patterns for WiMAX Networks with Adaptive Beamforming, *IEEE VTC Spring*, 2008
- [12] D. Catrein and R. Mathar, Gaussian random fields as a model for spatially correlated log-normal fading, *IEEE ATNAC*, 2008